Pseudo-Bayes Factors

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- Proposed approach: use a pseudo-BF based on a pseudo-likelihood L^{*}(ψ), which is a function of ψ only.

$$H_0: \psi \in \Psi_0$$
 against $H_1: \psi \in \Psi_1$,

with the Bayes Factor

$$BF = \frac{\int_{\Psi_0} \int_{\Lambda} L(\psi, \lambda) \pi_0(\lambda | \psi) \pi_0(\psi) \, d\lambda \, d\psi}{\int_{\Psi_1} \int_{\Lambda} L(\psi, \lambda) \pi_1(\lambda | \psi) \pi_1(\psi) \, d\lambda \, d\psi},$$

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- needs *p* dimensional integrations on (Ψ_k, Λ) .

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- Using $L^*(\psi)$ as a true likelihood, a posterior distribution of ψ can be considered

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where $\pi(\psi)$ is a suitable prior on ψ only.

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- Advantages in using $\pi^*(\psi|\mathbf{y})$:
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- In our approach we only need:
 - a pseudo-likelihood $L^*(\psi)$
 - priors $\pi_k(\psi)$ for $\psi \in \Psi_k$, k = 0, 1
- We then define the Pseudo-Bayes Factor:

$$BF^* = \frac{\int_{\Psi_0} L^*(\psi) \pi_0(\psi) \, d\psi}{\int_{\Psi_1} L^*(\psi) \pi_1(\psi) \, d\psi} = \frac{\int_{\Psi_0} \pi_0^*(\psi|\boldsymbol{y}) \, d\psi}{\int_{\Psi_1} \pi_1^*(\psi|\boldsymbol{y}) \, d\psi},$$

which needs only 1 dimensional integrations over Ψ_k .

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- To evaluate the behavior of *BF*^{*} we compare it versus *BF* in terms of the corresponding the Frequentist Risks, *R*^{*}(ψ, λ) and *R*(ψ, λ), where

$$\begin{array}{lll} R^*(\psi,\lambda) &=& \mathsf{Pr}(BF^*(\boldsymbol{y}) < 1|H_0) + \mathsf{Pr}(BF^*(\boldsymbol{y}) > 1|H_1) \\ R(\psi,\lambda) &=& \mathsf{Pr}(BF(\boldsymbol{y}) < 1|H_0) + \mathsf{Pr}(BF(\boldsymbol{y}) > 1|H_1) \end{array}$$

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- In the next examples we assume a favorable scenario for *BF*: $\pi_k(\lambda|\psi) = \pi(\lambda)$ for k = 0, 1, with $\pi(\lambda)$ concentrated on the true values.
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- In the next examples we assume a favorable scenario for *BF*: $\pi_k(\lambda|\psi) = \pi(\lambda)$ for k = 0, 1, with $\pi(\lambda)$ concentrated on the true values.
- Both R^{*}(ψ, λ) and R(ψ, λ) are approximated by Monte Carlo simulations.

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• We mainly focus on the Modified Profile Likelihood (Severini, 2000)

$$L_{mp}(\psi) = L_p(\psi) \frac{|\dot{j}_{\lambda\lambda}(\psi, \hat{\lambda}_{\psi})|^{1/2}}{|I(\psi, \hat{\lambda}_{\psi}; \hat{\theta})|} ,$$

where $L_{\rho}(\psi) = L(\psi, \hat{\lambda}_{\psi})$, $\hat{\theta}$ is the MLE for θ , $\hat{\lambda}_{\psi}$ is the conditional MLE of λ and

$$I(\psi, \lambda; \theta_0) = E_{\theta_0}(\ell_{\lambda}(\psi, \lambda)\ell_{\lambda}(\psi_0, \lambda_0)^{\mathsf{T}}),$$

with $\theta_0 = (\psi_0, \lambda_0)$ and $\ell_{\lambda}(\psi, \lambda) = \partial \ell(\psi, \lambda) / \partial \lambda$.

Example 1. Stress-strength model (Kotz et al., 2003)

Let X ~ Exp(α) and Y ~ Exp(β) be random variables. Interest is on

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Given random samples of size *n* and *m* from from X and Y respectively, we test H₀ : ψ < 1/2 against H₁ : ψ > 1/2, ∀λ > 0 assuming the following priors

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$$\pi_1(\psi) = U(1/2, 1)$$

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$$\pi(\lambda) = Gamma(1, 1)$$

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• The MPL is
$$L_{mp}(\psi) = -(n+m-2)\log\left(s_x + s_y \frac{1-\psi}{\psi}\right) + m\log\frac{1-\psi}{\psi}$$
,
where $s_x = \sum x_i$ and $s_y = \sum y_i$.

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- Both *BF** and *BF* are evaluated numerically.

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$$R^*$$
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 $\psi = 0.4$
 10%
 33%
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 $\psi = 0.6$
 29%
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 Table: Values of risks (with $n = m = 5$).

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- However, R is much more sensitive to λ than R^* .

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BF Risk - BFP Risk, n=2

BF Risk - BFP Risk, n=5









Figure: Values of $R(\psi, \lambda) - R^*(\psi, \lambda)$ for n = m = 2, 5, 10, 20.

Cabras et al. (CAG-PAD)

$$L(\beta) = \exp\left\{\sum_{i=1}^{n} y_i \sum_{j=1}^{p} \beta_j x_{ij} - \sum_{i=1}^{n} \log\left(1 + e^{\sum_{j=1}^{p} \beta_j x_{ij}}\right)\right\}$$

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with $\beta = (\beta_1, \dots, \beta_p)$ unknown regression coefficient and x_{ij} fixed constants, $i = 1, \dots, n$ and $j = 1, \dots, p$.

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 - $\pi(\lambda) = \pi(\beta_1) \cdot \ldots \cdot \pi(\beta_{p-1}), \quad \pi(\beta_j) = N(0, 10^{12}), \quad j = 1, 2, \ldots, p-1.$

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- Assume that ψ is the coefficient of the effect of the variable urea concentration.
- The weights of evidence (Good, 1985), W = log BF and W* = log BF*, are

W and W^*

<i>BF</i> based on $L(\beta)$	4.2	
BF^* based on $L_{mp}(\psi)$	4.2	

According to the Jeffreys' scale, the evidence is substantial in favor of positive effect of urea concentration.

Example 2. Logistic Regression - *p* **large.**

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Example 2. Logistic Regression - *p* **large.**

In order to assess the behaviour of BF^* , with respect to large *p*, we evaluate the corresponding W^* in 1000 data sets with p = 20 coefficients, with 5 positive, 5 negative and 10 zero coefficients.



Figure: Empirical mean W^* for the sign of 20 coefficients. (Horizontal dashed lines are the levels of strong evidence).

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Pseudo-Bayes Factors

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