

Bayesian Flexible Modelling of Mixed Logit Models

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Outline

- Random utility models and multinomial logit model
- Mixed logit models and the problem of choosing the mixing distribution
- A possible solution based on a discrete mixture of normal distributions
- An application to public transport demand data
- Conclusions

Random utility models

- Decision maker n chooses the alternative i , which maximises his own utility, among J possible alternatives:

$$U_{ni} > U_{nj} \quad \forall j \neq i$$

- The utility function can be written as:

$$U_{ni} = V_{ni} + \varepsilon_{ni}$$

- Thus, the probability that subject n chooses alternative i is:

$$P_n(i) = P(U_{ni} > U_{nj}) = P(\varepsilon_{nj} - \varepsilon_{ni} < V_{ni} - V_{nj}) \quad \forall j \neq i$$

Multinomial logit models (MNL)

- Assuming $\varepsilon_{ni} \sim$ i.i.d. *Extreme Value*, for all n, i

$$P_n(i) = \frac{e^{V_{ni}}}{\sum_{j=1}^J e^{V_{nj}}}$$

- If the utility function is linear:

$$P_n(i) = \frac{e^{\beta'x_{ni}}}{\sum_{j=1}^J e^{\beta'x_{nj}}}$$

Power and limitations of MNL

- ✓ Closed form formula for the choice probabilities
- ✗ Unobserved factors are uncorrelated over alternatives
- ✗ Independence from Irrelevant Alternatives (IIA)
- ✗ Proportional substitution between alternatives
- ✗ Temporal independence of unobserved factors (panel data)
- ✗ Homogeneous preference structure

Random parameter logit models

- ✿ The mixed MNL (MMNL) assumes the vector β having density $f(\beta|\theta)$,

$$P_n(i) = \int \frac{e^{\beta'x_{ni}}}{\sum_{j=1}^J e^{\beta'x_{nj}}} f(\beta|\theta) d\beta$$

- ✿ The latent class MNL assumes that the population is composed of C homogeneous groups

$$P_n(i) = \sum_{c=1}^C \left(\frac{e^{\beta_c'x_{ni}}}{\sum_{j=1}^J e^{\beta_c'x_{nj}}} \right) s_c$$

MMNL specification issues

- ✿ Selecting the parameters that are to be random parameters
 - Lagrange multiplier test (McFadden & Train, 2000)
- ✿ Selecting the distribution of the random parameters
 - some parameter needs to have a specific sign
 - most distributions have long tails
- ▶ Hensher & Greene (2003) suggest empirical methods to help in the choice
- ▶ Hess *et al.* (2005) compare different distributions using simulated and real data
- ▶ Fosgerau and Hess (2009) proposed two approaches:
 - 1) improving on the flexibility of a base distribution through Legendre polynomials
 - 2) making use of a semi-parametric mixing distribution consisting of a discrete mixture of normal distributions

The proposed approach

- The utility of person n from alternative i in period t is:

$$U_{nit} = \alpha' w_{nit} + \beta'_n x_{nit} + \varepsilon_{nit}$$

- We assume:

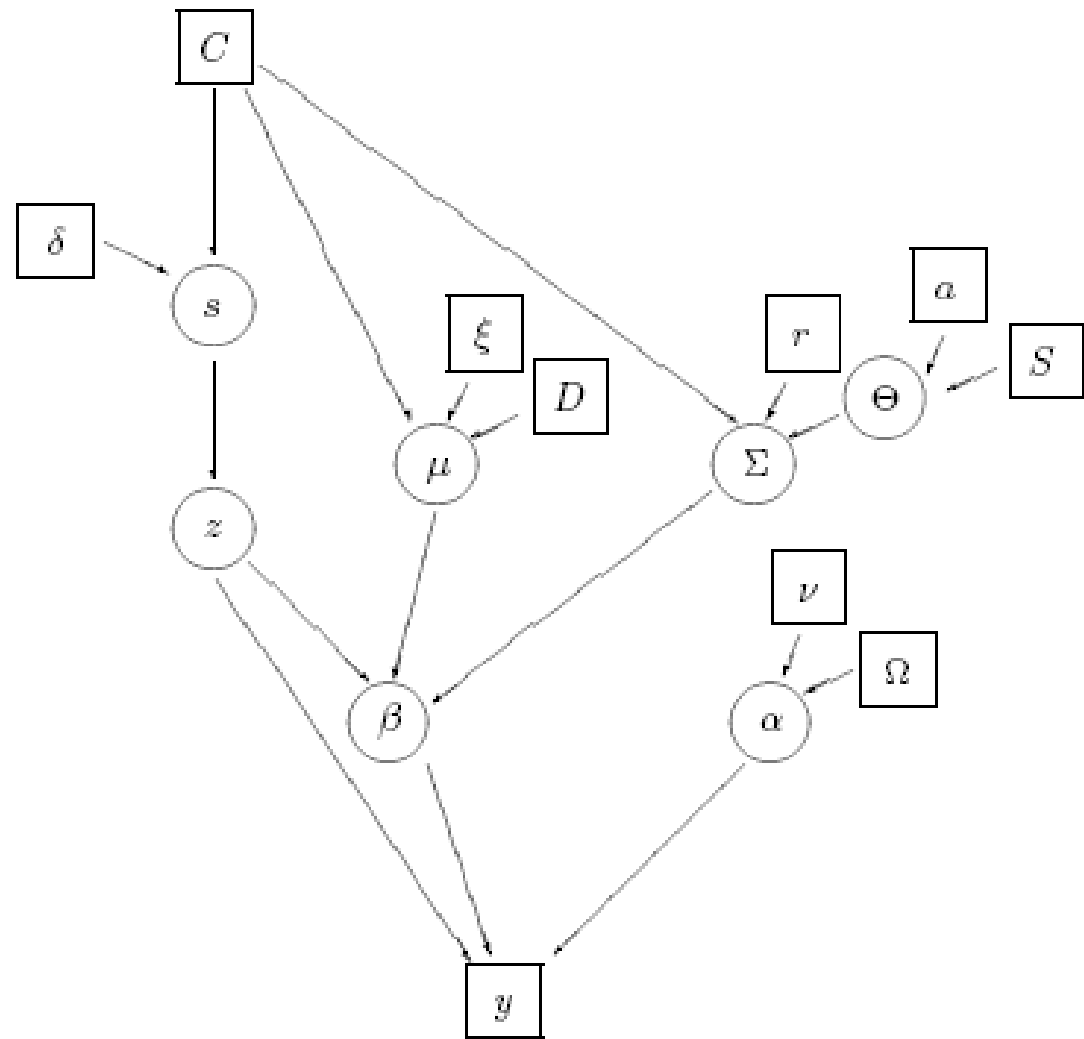
$$\beta_n \mid \mu, \Sigma \sim \sum_{c=1}^C s_c \phi(\cdot \mid \mu_c, \Sigma_c)$$

- Or, analogously, introducing latent variables $z = \{z_1, \dots, z_N\}$:

$$\beta_n \mid z, \mu, \Sigma \sim \phi(\cdot \mid \mu_{z_n}, \Sigma_{z_n})$$

Prior distributions on the parameters

- $C = 2;3$
- $(s_1, \dots, s_C) \sim D(\delta, \dots, \delta)$
- $z_n \sim p(z_n = c) = s_c$
- $\mu_c \sim N(\xi, D)$
- $\Sigma_c \sim IW(r, \Theta^{-1})$
- $\Theta \sim IW(a, S)$
- $\alpha \sim N(\nu, \Omega)$



Bayesian Inference

- We use MCMC to sample from the posterior joint distribution of the parameters
 - ➔ update s , z , μ , Θ and Σ through Gibbs steps
 - ➔ update α and β through Metropolis steps
- From the sample $(s^{(m)}, z^{(m)}, \mu^{(m)}, \Theta^{(m)}, \Sigma^{(m)}, \alpha^{(m)}, \beta^{(m)})$, for $m = 1, \dots, M$, we estimate quantities of interest, i.e.:
 - ➔ estimated individual-level taste parameters

$$\hat{\beta}_n = \frac{1}{M} \sum_{m=1}^M \beta_n^{(m)}$$



An application to public transport demand

- Aim of the study: analyze attributes of local public transport in [Urbino](#) (Italy) and investigate possible intervention to improve the service
- Attributes considered:
 - cost of monthly ticket (5 levels)
 - headway (5 levels)
 - first and last run (5 levels)
 - real time information displays (2 levels)
 - bus shelters (2 levels)
- Questionnaire: 15 choice exercises of which
 - 11 random
 - 2 aimed at testing the quality of the answers
 - 2 aimed at testing preference stability
- Data set: 50 respondents took part in the study, providing a data set of 750 observations

An application to public transport demand

If you were to use the public transport Mercatale-Sogesta and the service would have the following characteristics, which alternative would you choose?			
Prize 15.40 €	Prize 20.80 €	Prize 12.80 €	None of these alternatives
Headway: 30 minutes	Headway: 15 minutes	Headway: 60 minutes	
First and last run 07:15 -- 01:00	First and last run 08:15 -- 24:00	First and last run 07:15 -- 01:00	
No real time information displays	Real time information displays	Real time information displays	
Bus shelters at Mercatale & Sogesta	Bus shelters at Mercatale & Sogesta	Bus shelters at Mercatale	
A	B	C	

Results

-  Lagrange multiplier test to decide which parameters are to be random
 - headway
 - first and last run
-  100,000 sweeps of the MCMC algorithm, with a burn-in of 50,000 sweeps

Parameter	MMNL		MOD (2 components)		MOD (3 components)	
	Est.	Std. dev.	Est.	Std. dev.	Est.	Std. dev.
α_{cost}	-0.2671	0.0336	-0.2751	0.0346	-0.2764	0.0349
α_{displays}	-0.0452	0.1753	-0.0190	0.1678	-0.0242	0.1828
α_{shelters}	0.3258	0.1771	0.3167	0.1835	0.3320	0.1747
μ_{headway}	-0.1039	0.0247	-0.0673	0.0806	-0.0485	0.1525
			-0.1116	0.0254	-0.0750	0.0864
					-0.1127	0.0265
$\mu_{\text{run time}}$	0.5322	0.0537	0.0403	0.1454	0.0180	0.1926
			0.6561	0.0798	0.0863	0.1697
					0.6840	0.0867
s	1.0000	0.0000	0.1599	0.0891	0.0503	0.0398
			0.8401	0.0891	0.1623	0.0818
					0.7873	0.0957

Table 1. *Posterior mean estimates of relevant parameters.*

Results

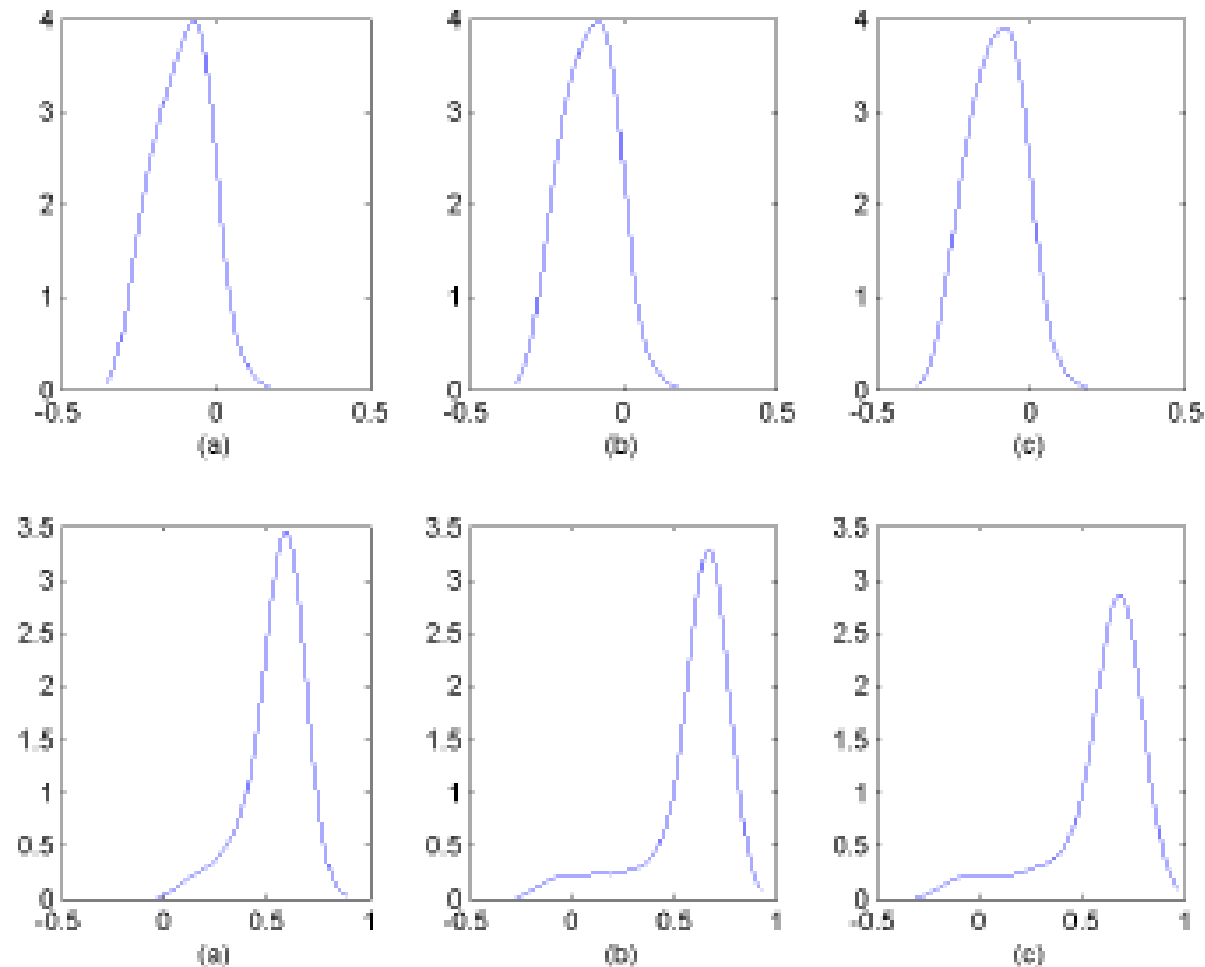


Fig. 1. Estimated marginal posterior densities for the random parameters β_{headway} (upper panel) and $\beta_{\text{run time}}$ (lower panel) under a) the MMNL model, b) the MOD model with 2 components, c) the MOD model with 3 components.

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